

# Statistical Information Theory and Geometry for SAR Image Analysis

Recent results and research topics

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- Connecting SAR (*Synthetic Aperture Radar*) image processing and analysis with statistical models
- Commenting upon seemingly disjoint problems
- Showing that they can be tackled successfully within the same framework
- Discussing the use of Information-Theoretic Statistically-Based tools

# SUMMARY

SAR Imagery

The Multiplicative Model

Assessment of Despeckling Filters

Model for Intensity Data

Model for Polarimetric Data

Distances between Kennaugh projections

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# WHAT IS SAR?



## BASIC CHARACTERISTICS

- Synthetic Aperture Radar (SAR) and Polarimetric SAR (PolSAR) sensors have been successfully used in remote sensing.
  - data can be captured independently of weather conditions, because the sensor is active,
  - they measure mostly geometry and dielectric constant, as they work in the microwaves region of the spectrum,
  - depending on the sensor and target characteristics, the signal penetrates through soil, canopies etc.
- SAR and PolSAR systems can provide images with high spacial resolution but contaminated by an interference pattern, called **speckle**.

# SAR IMAGE



# GENERAL DISCUSSION

- ☞ **Speckle** imposes a peculiar behavior to the SAR return:
  - a good model for the return is multiplicative, rather than additive, and
  - intensities follow non-Gaussian distributions.
- ☞ This makes SAR image analysis challenging because standard methods for image processing
  1. assume additivity  $Z = X + Y$ , and
  2. use tools derived from the Gaussian assumption.
- ☞ Observations are organized as a complex-valued Hermitian positive definite matrix in each position of a fully PolSAR image.
  - Thus **specialized** tools are required.

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# THE MULTIPLICATIVE MODEL

A well accepted model for the return  $Z$  in each pixel is the Multiplicative Model:

$$Z = XY,$$

where  $X$  is the backscatter, and  $Y$  is the speckle.

The physics of the imaging allows to assume that  $Y$  follows a Gamma distribution with unitary mean and shape parameter  $L \geq 1$ , the number of looks. We denote this  $Y \sim \Gamma(1, L)$ .

We are interested in the backscatter  $X$ , which can be constant or random but always positive.

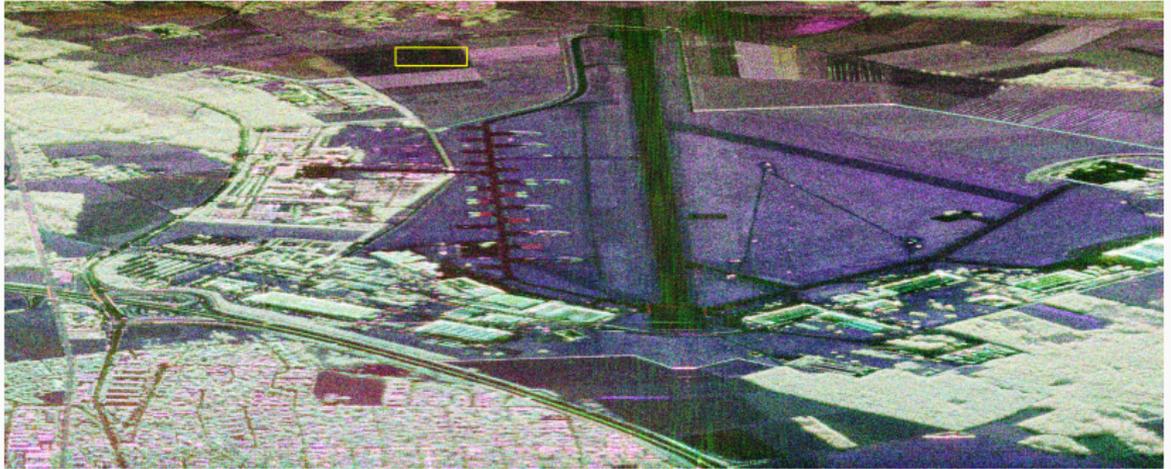
# IS THIS MODEL CORRECT?

Box et al. (2005)

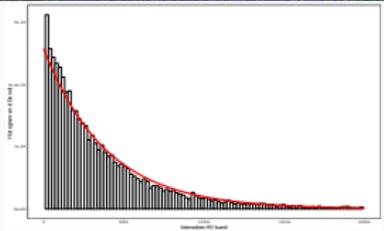
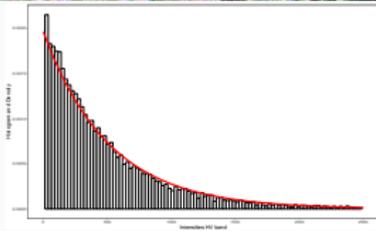
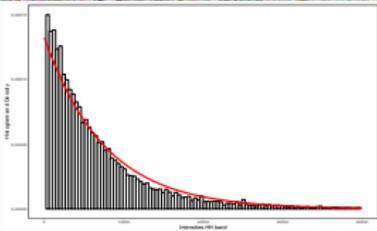
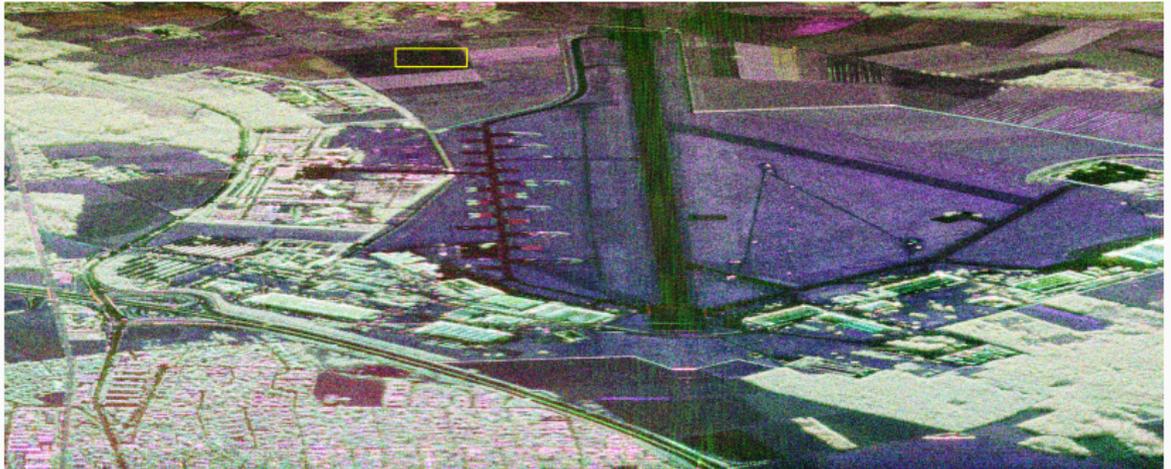
The most that can be expected from any model is that it can supply an useful approximation to reality:

***All models are wrong; some models are useful.***

# IS THIS MODEL CORRECT?



# IS THIS MODEL CORRECT?



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## ASSESSMENT OF DESPECKLING FILTERS

There are many despeckling techniques, as well as measures of their performance.

Assuming the multiplicative model, the observed image  $Z$  is the product of two independent fields: the backscatter  $X$  and the speckle  $Y$ .

The result of any speckle filter is  $\hat{X}$ , an estimator of the backscatter  $X$ , based solely on the observed data  $Z$ .

An ideal estimator would be the one for which the ratio  $I = Z/\hat{X}$  is only speckle: a collection of independent identically distributed samples from Gamma variates.

We, then, assess the quality of a filter by the **closeness** of the ratio image  $I$  to the hypothesis that it is adherent to the statistical properties of pure speckle.

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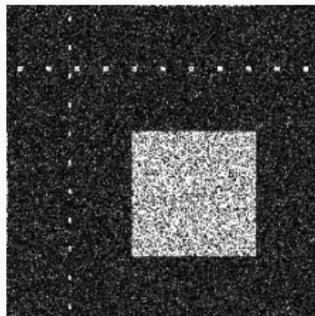
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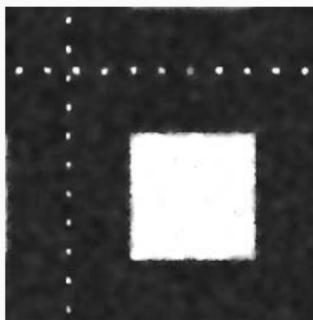
We propose an evaluation of the quality of the remaining speckle based on two components:

**First-order component:** one term for mean preservation, and another for preservation of the equivalent number of looks.

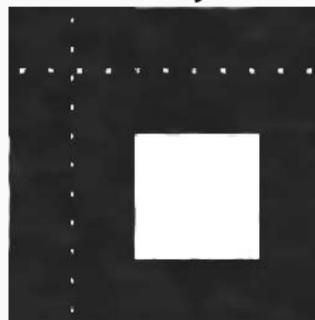
**Second-order component:** a measure of the remaining geometrical content within the ratio image.



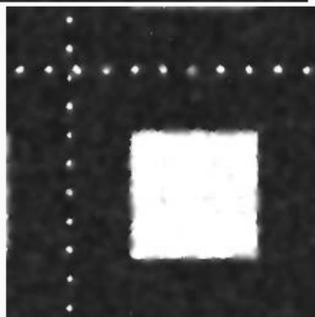
Noisy



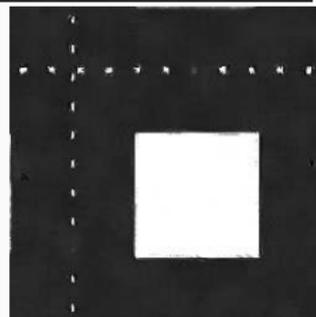
Lee



FANS



SRAD



PPB

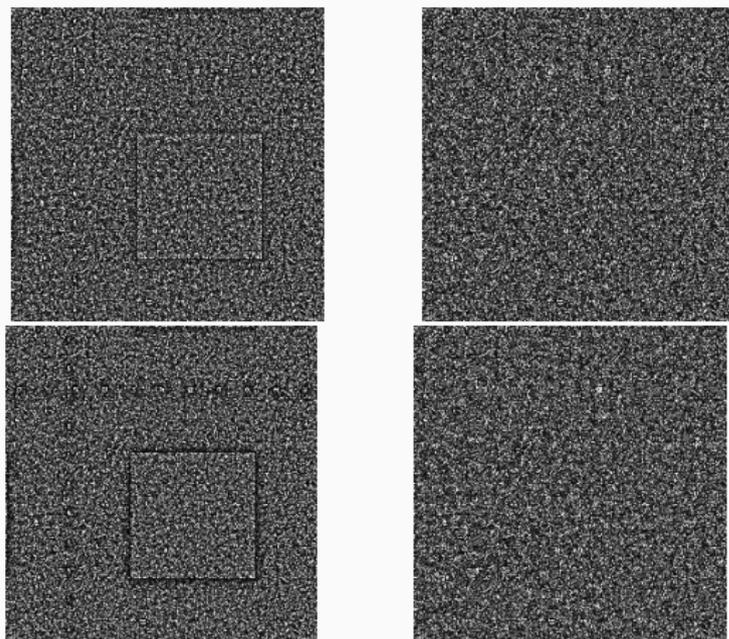


Figure 1: Top: E-Lee and FANS. Bottom: SRAD and PPB.

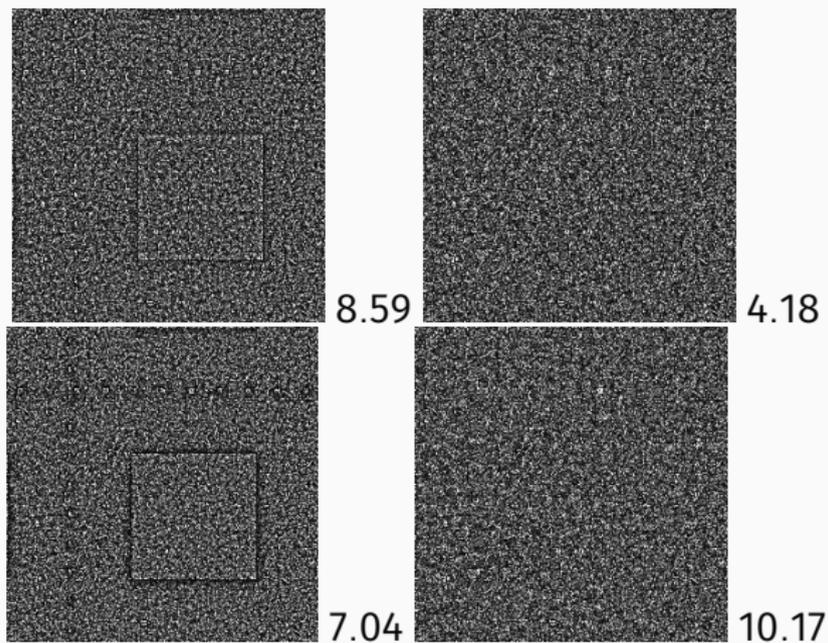


Figure 1: Top: E-Lee and FANS. Bottom: SRAD and PPB.

The rationale behind this proposal is measuring the distance between two objects, namely

- a theoretical distribution, and
- observed data.

This was done in an empirical manner, and provided very good results; cf. Gomez et al. (2017*b*, 2019).

## A GENERAL APPROACH

This idea of comparing samples, or samples and models, has led to interesting tools.

First, we will see a model with respect to which such comparisons will be made. Then we will see measures based on Information Theory and Information Geometry. Finally, we will see a multivariate model for PolSAR data, and several applications of these techniques.

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# THE $G^0$ RETURN MODEL

We call  $Y \sim \Gamma(1, L)$  **speckle**, and  $X = \sigma^2$  **backscatter**, and we assume these random variables are independent.

The **return** is  $Z = XY$ .

When the mean backscatter  $\sigma^2$  fluctuates, the return  $\sigma^2 Y$  is no longer Gamma-distributed.

The two most important stochastic models for the backscatter are:

- A Gamma random variable
- A Reciprocal Gamma random variable

# THE $\mathcal{G}^0$ RETURN MODEL

Assuming  $X \sim \Gamma^{-1}(\alpha, \gamma)$  and  $Y \sim \Gamma(1, L)$ , one obtains the  $\mathcal{G}^0$  distribution for the return  $Z = XY$ , which is characterized by the density

$$f_Z(z; \alpha, \gamma, L) = \frac{L^L \Gamma(L - \alpha)}{\gamma^\alpha \Gamma(L) \Gamma(-\alpha)} \frac{z^{L-1}}{(\gamma + Lz)^{L-\alpha}}, \quad (1)$$

where  $\alpha < 0$ , and  $\gamma, z > 0$ .

The parameters describe the texture (the smaller the value of  $\alpha$  is, the less textured the region is), and the scale ( $\gamma$ ).

This model was proposed by Frery et al. (1997).

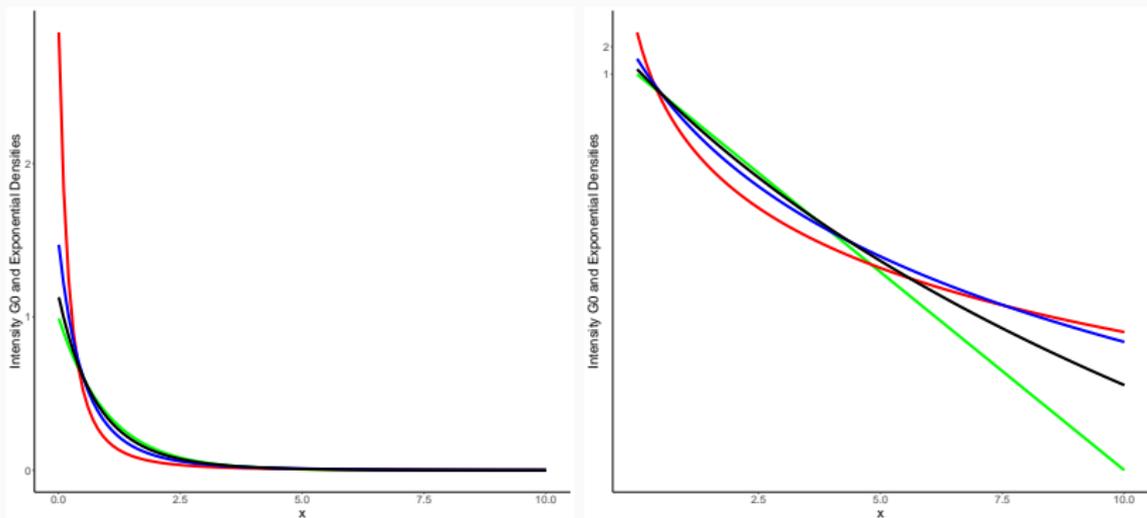


Figure 2: Densities in linear and semilogarithmic scale of the  $E(1)$  (green) and  $\mathcal{G}^0$  distributions with unitary mean and  $\alpha \in \{-8, -3, -1.5\}$  in black, blue and red, resp.

# ADVANTAGES AND DISADVANTAGES

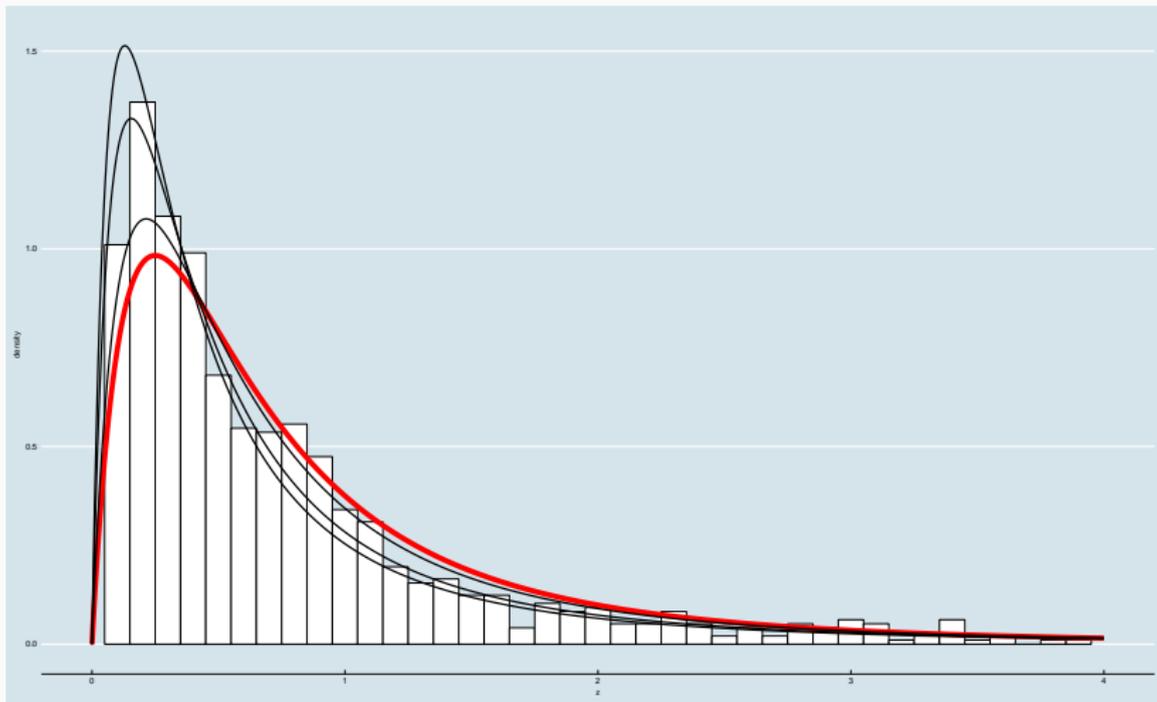
## The Good

- More flexible than the  $\Gamma$  model
- It has the  $\Gamma$  model as particular case
- It relates to the Fisher-Snedekor distribution
- Differently from the  $\mathcal{K}$  law, it does not involve Bessel functions

## The Bad

- Its likelihood function may become flat
- Its moments estimators may have no solution
- Those estimators are susceptible to contamination

# PARAMETER ESTIMATION



# PARAMETER ESTIMATION

Gambini et al. (2015) proposed estimating by searching in  $\Theta$  for the point  $\hat{\theta}$  which minimizes a distance between the evidence (data) and the model:

$$\hat{\theta} = \arg \min_{\theta \in \Theta} d(\tilde{H}(\mathbf{z}), f(\mathbf{z}; \theta)),$$

where  $\mathbf{z}$  are the data, and  $\tilde{H}$  is an smoothed version of the histogram. Also,

$$d(f_1, f_2) = \int \frac{(f_1 - f_2)^2}{f_1 + f_2},$$

is the Triangular (or symmetric  $\chi^2$ ) distance.

It works quite well!

## WHY SUCH A DISTANCE?

We chose the Triangular distance, because of its numerical simplicity and stability, within a family of divergences between members of the same law:

*h*- $\phi$  divergences

$$D_{\phi}^h(\theta_1, \theta_2) = h\left(\int \phi\left(\frac{f(\theta_1)}{f(\theta_2)}\right)f(\theta_2)\right),$$

where  $h$  is strictly increasing with  $h(0) = 0$ , and  $\phi$  is convex with mild smoothness properties. Then,

*h*- $\phi$  distances

$$d_{\phi}^h(\theta_1, \theta_2) = \frac{1}{2}(D_{\phi}^h(\theta_1, \theta_2) + D_{\phi}^h(\theta_2, \theta_1)).$$

Suitable choices of  $h$  and  $\phi$  lead to, among others, the Bhattacharya, Rényi, Triangular, Harmonic,  $\chi^2$ , Kullback-Leibler, Hellinger, Havrda-Charvát, and Sharma-Mittal distances.

## ANOTHER FAMILY OF DISTANCES

Consider the model  $\mathcal{D}(\theta)$ , two parameters  $\theta^1, \theta^2 \in \Theta$ , and let  $t$  be the parameter of a curve  $\theta(t) \in \Theta$  which joins  $\theta^1 = \theta(t_1)$  and  $\theta^2 = \theta(t_2)$ .

The (Shannon) geodesic distance between the models is given by:

$$s(\theta^1, \theta^2) = \left| \int_{t_1}^{t_2} \sqrt{\sum_{i,j=1}^r -E\left(\frac{\partial^2}{\partial\theta_i\partial\theta_j} \ln f(z | \theta)\right) \frac{d\theta_i}{dt} \frac{d\theta_j}{dt}} dt \right|. \quad (2)$$

This expression can be connected with the entropy of  $\mathcal{D}(\theta)$ , and with generalized entropies:

$$H_{\phi}^h(\theta) = h\left(\int \phi(f)f\right),$$

leading to  $h$ - $\phi$  geodesic distances.

This is **tough** to obtain in most cases.

# WHAT DO WE HAVE SO FAR?

- It seems that the distance between the data and a model produces interesting results:
  - **measuring** it quantifies the quality of filters;
  - **minimizing** it yields robust estimators.
- But... is there a notion of **close** and **far**? **Not yet.**
- Can we compare such distances? **No!**

When indexed by MLE, under  $H_0 : \theta_1 = \theta_2$  holds that

$$\frac{N_1 N_2}{(N_1 + N_2) h'(0) \phi''(1)} d_{\phi}^h(\widehat{\theta}_1, \widehat{\theta}_2)$$

follows asymptotically a  $\chi_p^2$  distribution, where

- $p$  is the dimension of  $\theta_i$ ,
- $N_1$  and  $N_2$  are the sample sizes used to compute the MLEs  $\widehat{\theta}_1$  and  $\widehat{\theta}_2$ ,
- $N_1, N_2 \rightarrow \infty$  at comparable rate.

For the sake of simplicity, but without loss of generality, assume we have two samples of the same size  $N$ . When indexed by MLE, under  $H_0 : \theta_1 = \theta_2$  holds that

$$N \frac{(H(\widehat{\theta}_1) - H(\widehat{\theta}_2))^2}{\sigma^2(\widehat{\theta}_1) + \sigma^2(\widehat{\theta}_2)}$$

follows asymptotically a  $\chi_{p-1}^2$  distribution, where  $p$  is the dimension of  $\theta_i$ , and  $N \rightarrow \infty$ .

This can be generalized to many samples of different sizes.

# WHAT DO WE HAVE NOW?

## Test Statistics based on IT

- Each  $h$ - $\phi$  divergence can be turned into a test statistic.
- Each  $h$ - $\phi$  entropy can be turned into a test statistic.

## Test Statistics based on IG

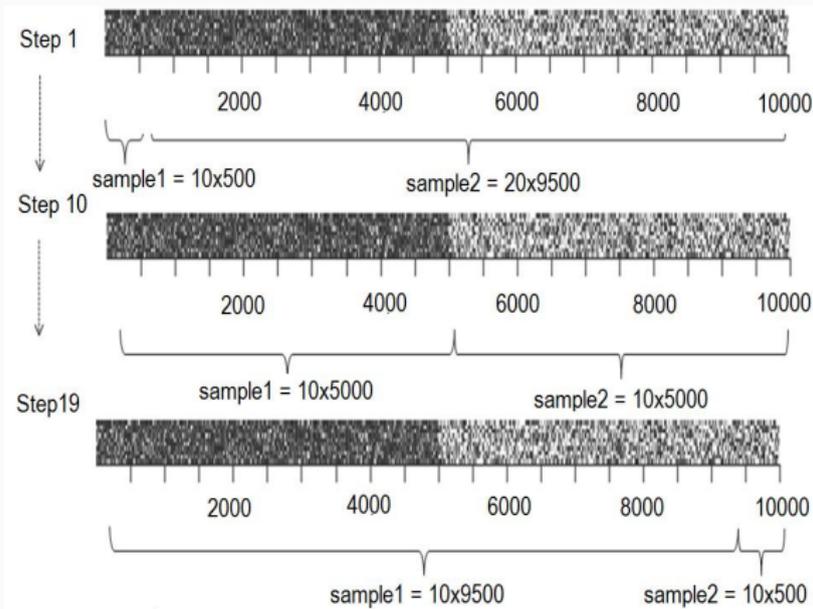
Each geodesic distance can be turned into a test statistic.

All these test statistics have known asymptotic distribution, so they are **interpretable** and **comparable**.

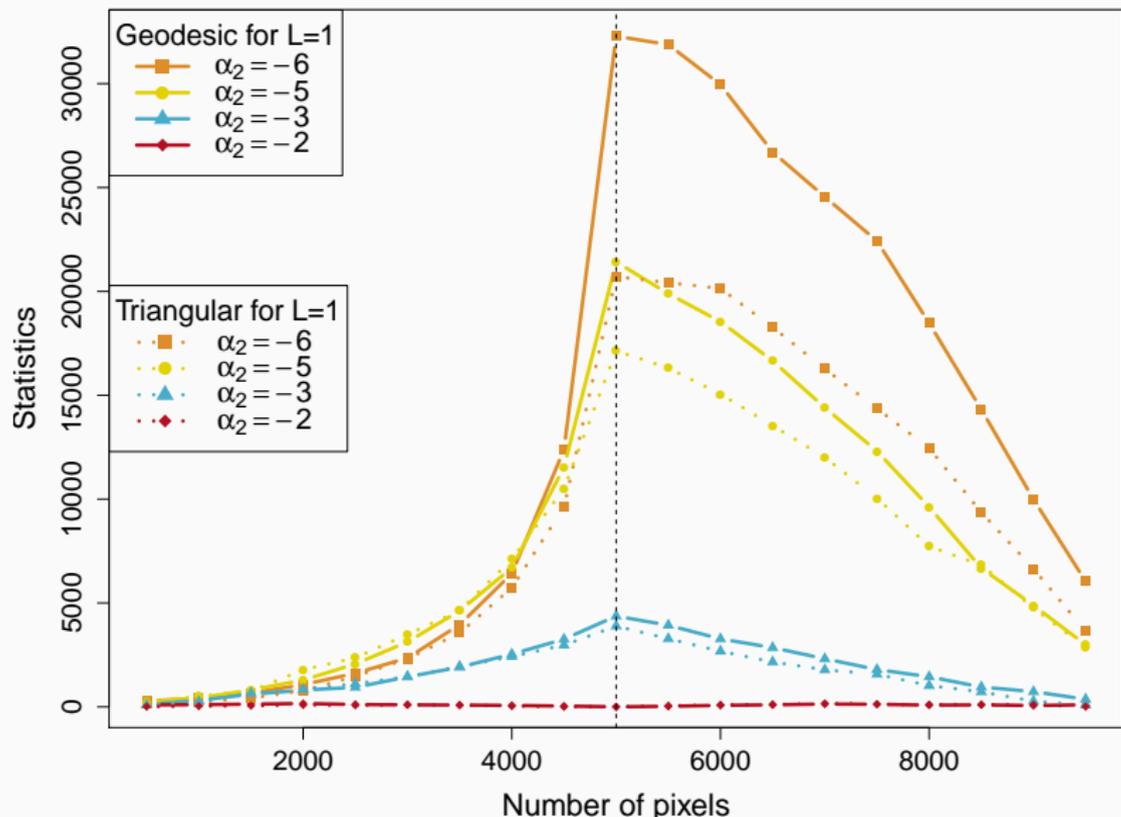
Moreover, they exhibit excellent finite-size behavior: they have good size and power even with very small samples.

# EDGE DETECTION WITH GD UNDER THE $\mathcal{G}^0$ LAW

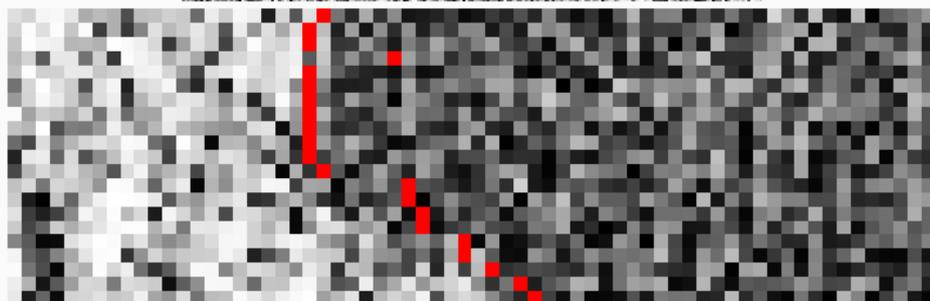
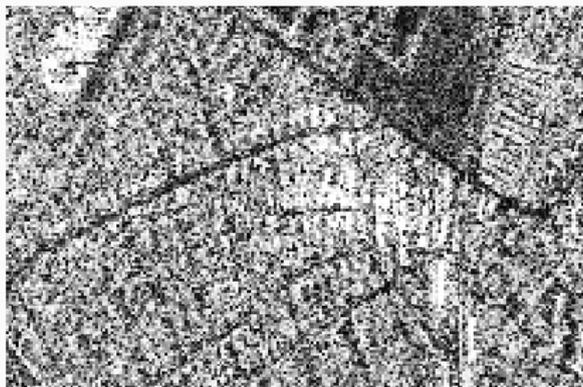
Naranjo-Torres et al. (2017) obtained expressions for the GD between  $\mathcal{G}^0$  models and proposed a line search algorithm for edge detection.



# EDGE DETECTION WITH GD



# EDGE DETECTION WITH GD



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# WHAT IS POLSAR?



# THE COMPLEX WISHART DISTRIBUTION

Analogously to the  $\Gamma(\sigma^2, L)$  for intensity data, the Complex Wishart distribution is a widely accepted model of fully polarimetric data. It is characterized by the density

$$f_{\mathbf{Z}}(\mathbf{z}; \Sigma, L) = \frac{L^p |\mathbf{z}|^{L-p}}{|\Sigma|^L \Gamma_p(L)} \exp[-L \operatorname{tr}(\Sigma^{-1} \mathbf{z})],$$

where  $\Gamma_p(L) = \pi^{p(p-1)/2} \prod_{i=0}^{p-1} \Gamma(L-i)$ ,  $L \geq p$ , and  $\operatorname{tr}(\cdot)$  is the trace operator. We denote it by  $\mathbf{Z} \sim \mathcal{W}(\Sigma, L)$ .

This distribution satisfies  $E\{\mathbf{Z}\} = \Sigma$ , which is a Hermitian positive definite matrix (Anfinsen et al., 2009). In practice,  $L$  is treated as a parameter and must be estimated, usually for the whole image.

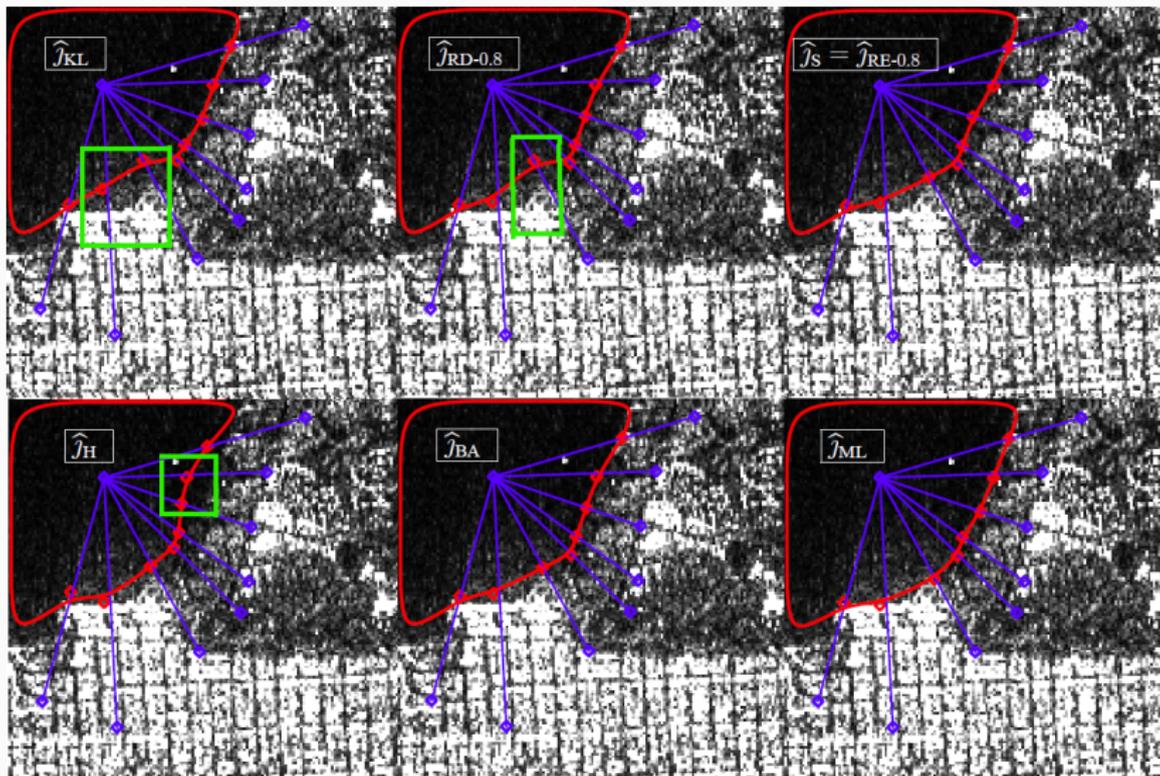
## PROPERTIES OF THE $\mathcal{W}$ DISTRIBUTION

- It is a distribution for complex Hermitian matrices.
- The marginal distribution of each diagonal element is a  $\Gamma(\sigma^2, L)$  law.
- The joint distribution of the diagonal elements was derived by Hagedorn et al. (2006).
- Other important marginal laws were obtained by Lee et al. (1994).

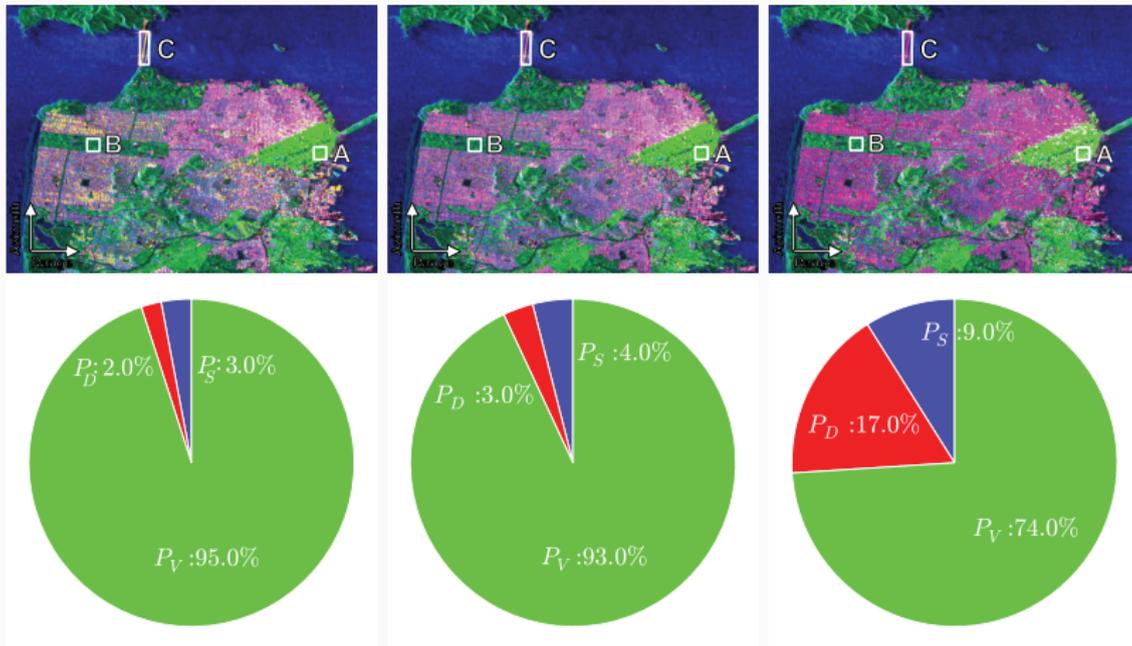
Frery et al. (2014) computed several of these tests under the  $\mathcal{W}$  model:

- Relatively simple expressions
- They all rely only on the determinant and the inverse of  $\Sigma$
- Good size and power even for very small (4) samples

# EDGE DETECTION STOCHASTIC DISTANCES



# SPECTRAL DECOMPOSITION BY ROTATION



The screenshot displays the POLARPRO software interface. At the top, the logo for 'esa' (European Space Agency) and 'POLARPRO' is visible, along with the tagline 'The Polarimetric SAR Data Processing and Educational Tool'. Below the logo is a menu bar with options: Environment, Import, Convert, Process, Display, Calibration, Utilities, Tools, Configuration, Education, and Help.

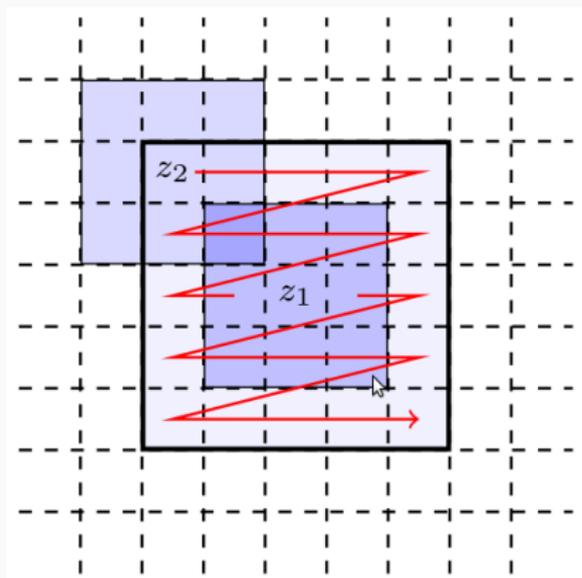
The main workspace shows a Windows desktop with various icons. A dialog box titled 'Data Processing: Polarimetric Decomposition' is open in the center. It contains the following fields and options:

- Input Directory:** F:\SAR Data\NRSARF\_SanFrancisco\_LEE\T3
- Output Directory:** F:\SAR Data\NRSARF\_SanFrancisco\_LEE
- Range:** Init Row: 1, End Row: 300, Init Col: 1, End Col: 1024
- Method:** Bhattacharya & Frey 4 Component Decomposition T3
- Decomposition Options:**
  - Yamaguchi Decomposition
    - Four Component Decomposition (Original: Y40)
    - Four Component Decomposition with Rotation Transformation (Y4R or S4R)
  - Singh Decomposition
    - Four Component Decomposition with Special Unitary Transformation (G4U1 or G4U2)
  - Bhattacharya & Frey Decomposition
    - Four Component Decomposition with stochastic distance (Original: Y40)
- Volume Scattering Model (Automatic Estimation):**
  - With / Without Extended Volume Scattering Model (diagonal scattering)
- Target Generators (TgG):**  TgG1  TgG2  TgG3
- Window Size:** Row: 3, Col: 3
- Buttons:** Run, Stop, Exit

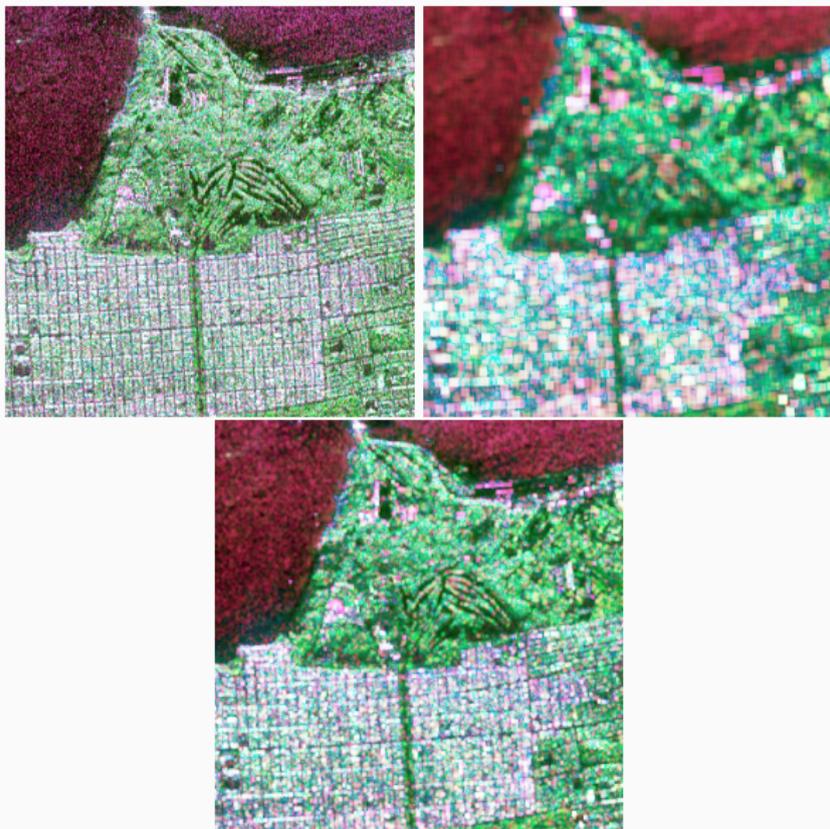
On the right side of the interface, a window titled '#2 R.Bhattacharya\_Frey\_Y40\_Dolbin\_Bhattacharya...' displays a SAR image with a color-coded decomposition. The image shows a dense urban or forested area with various colors representing different scattering mechanisms. Below this window, two smaller windows are visible: '#2 Scroll (0.25000)' and '#2 Zoom (4x)', showing zoomed-in views of the SAR image.

# NONLOCAL MEANS SPECKLE REDUCTION

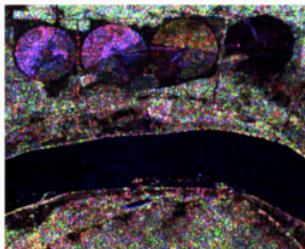
A large convolution mask is computed at each position, with weights inversely proportional to the dissimilarity between the point and the central value.



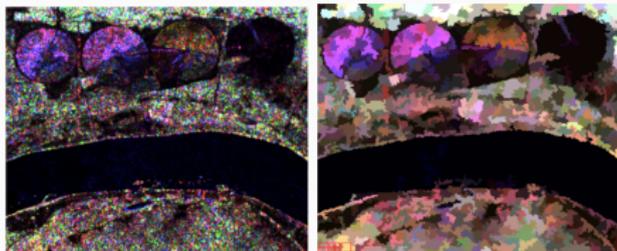
# NONLOCAL MEANS SPECKLE REDUCTION



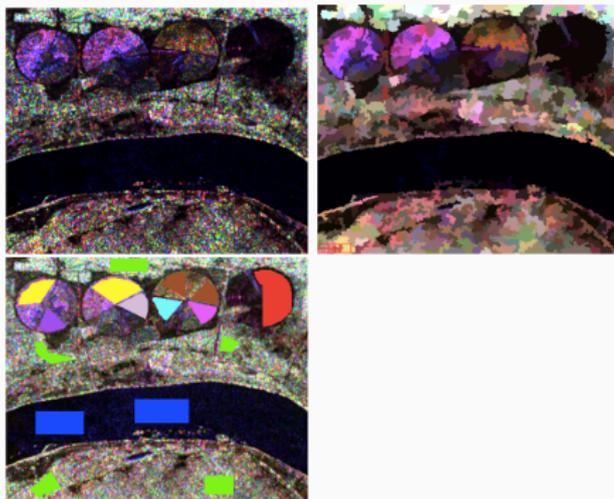
# FROM SEGMENTS TO CLASSES



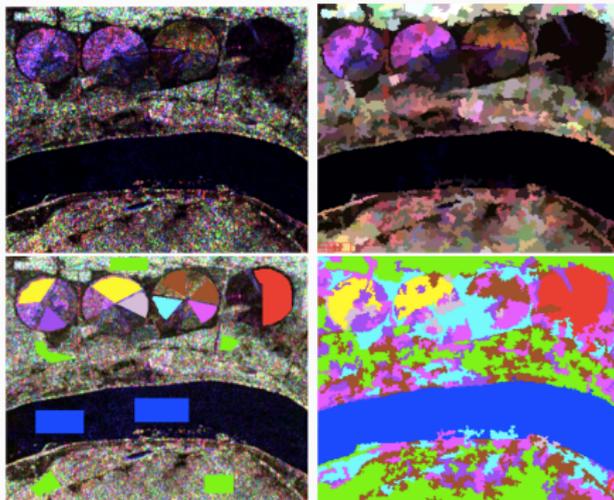
# FROM SEGMENTS TO CLASSES



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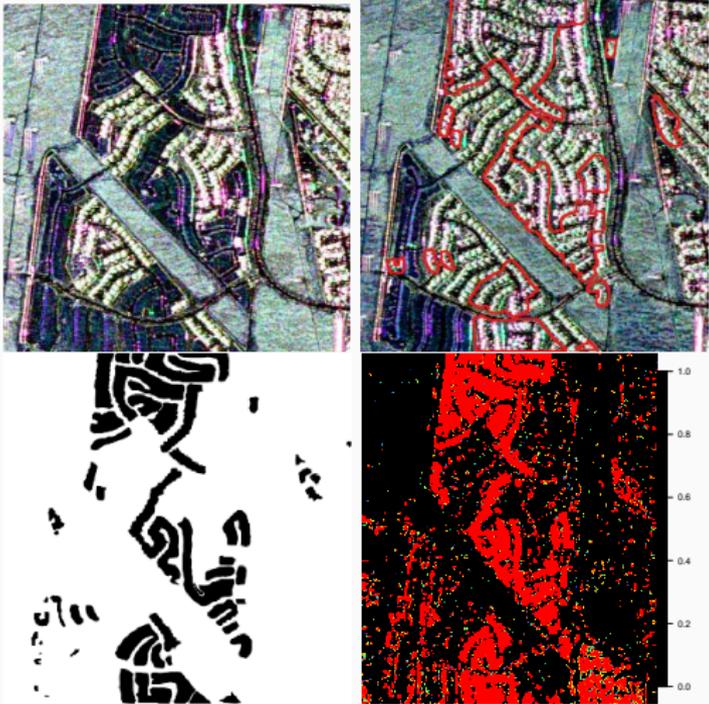
## DIFFERENCE OF ENTROPIES

It provides a powerful result for checking if at least one of the scenes under test has changed from the rest.

Being a test statistic, it is possible to compute the  $p$ -value of the no-change hypothesis at each position, rather than a binary mask.

This was applied to pairs of fully PolSAR images.

# CHANGE DETECTION



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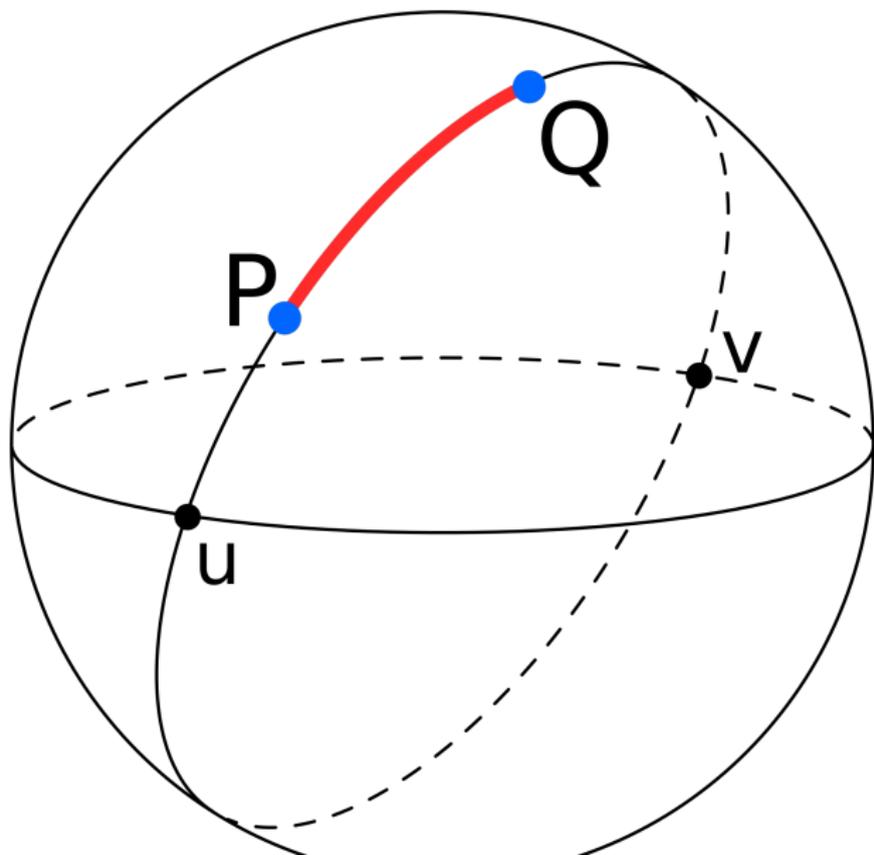
# THE FRONTIERS: KENNAUGH PROJECTION

Fully PolSAR data, and elementary backscatterers, can be projected onto the surface of a 16-dimensional sphere.

Then, distances become angles.

By measuring angles, we can classify, and decompose observations.

ONCE PROJECTED...



## CONCLUDING REMARKS

- Several interesting problems can be posed in the form of comparing samples and/or models.
- Information Theory and Information Geometry, along with Statistics, provide us with tools for solving such problems.
- Although the focus was on SAR/PolSAR, the basic ideas can be used whenever there is a model... even when there is no model!
- This is a fertile area for scientific research and technology development.

We only presented the geodesic distance for  $p$ -dimensional parameters, which is seldom feasible.

The alternative we worked with consisted in dealing with a one-dimensional parameter after scaling the data.

There is an another approach which allows *mixing* one-dimensional test statistics and using other geodesic metrics. This approach is closely related to generalized entropies.

- What happens with the convergence when we use non-ML estimators?
- What happens with the convergence when we use kernel estimators?
- Can we obtain other geodesic distances, stochastic distances, and difference of entropy, and their tests?
- Can we turn them into other useful techniques?
- What can we say about the statistics of distances in the Kennaugh space?

# REFERENCES I

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